NEWS-DRIVEN SYSTEMIC TAIL RISK AT HIGH FREQUENCY*

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INTRODUCTION

MOTIVATIONAL BACKGROUND

- Financial markets are often vulnerable to tail shocks driven by news.
 - → Some examples: the crash of "Black Monday" (October 19, 1987), the bailout rejection (September 29, 2008), flash crash (May 6, 2010), FOMC rate cut (December 11, 2007), the COVID-19 crash (March 16, 2020), ...
- When news-induced shocks hit, asset and portfolio managers are exposed to large financial losses associated with tail risks.
- The potential losses due to realized tail episodes may pose severe challenges:
 - \rightarrow Asset managers/investors/traders: When does the tail risk occur? How frequently? What triggers that? Is it diversifiable? Is it priced?
 - \rightarrow Policymakers/regulators: How "bad" is the financial distress? Is it a short-term market reaction or is it linked to policy decisions?

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INTRODUCTION

Related Literature

PRIOR RESEARCH

Idiosyncratic and systematic jump risks matter:
 Pelger (2020, JF), Chan et al. (2017, JEF), Bollerslev et al. (2013, 2008, JoE)

Tail risk is systematic and priced in the market: Andersen et al. (2020, *JBES*), Weller (2019, *RFS*), Van Oordt and Zhou (2016, *JFQA*), Bollerslev and Todorov (2014, *JoE*), Bollerslev et al. (2013, *JoE*), ...

Ongoing debate on the link between news announcements and jump risk: Lahaye et al. (2011, JAE), Amengual and Xiu (2018, JoE), Bajgrowicz et al. (2016, MS), ...

OUR FOCUS

Identification and financial implications of systemic tail risk: (risk that occurs when financial assets jump or crash together at the same time)

 \rightarrow Das and Uppal (2005, *JF*):

"Weak evidence" that systemic risk matters for international asset allocation

 \rightarrow Caporin et al. (2017, *JFE*):

"Strong evidence" that systemic risk matters, reveals return predictability

INTRODUCTION

This Paper: Contributions and Findings

(1) <u>Methodological</u>:

- We develop a new methodology to measure news-induced systemic tail risk:
 - $\rightarrow~$ Conditional testing and inference based on news release times
 - \rightarrow Exploiting time-varying jump intensity dynamics to capture tail risk
 - → Accurate identification (systemic volatility risk versus systemic jump risk)
 - \rightarrow Conservative bias control for spurious detection + bootstrap consistency

(2) <u>Empirical</u>:

- We use a panel of HF data on individual stocks and sector portfolios to study if U.S. monetary policy (FOMC) announcements lead to systemic tail risk:
 - \rightarrow We find strong evidence of "Fed-driven" systemic tail risk.
 - \rightarrow Left tail (systemic crash) risk occurs frequently over the business cycle.
 - \rightarrow Results hold for <u>both</u> individual stocks and sector (ETF) portfolio indices.
 - \rightarrow We identify which Fed events are systemically important as tail events.
 - \rightarrow We construct a simple proxy for systemic tail risk:
 - Helps explain the pre-FOMC announcement drift puzzle
 - Predicts intraday stock returns ahead of the upcoming Fed meeting
 - $\rightarrow\,$ No clear evidence that macro news creates HF systemic tail risk.

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The General Form

JUMP-DIFFUSION WITH TIME-VARYING INTENSITY

Log-prices $X := [X_1, \ldots, X_N]'$ of N assets:

$$dX_{i,t} = b_{i,t}dt + \sigma_{i,t}dW_{i,t} + \xi_{i,t}dJ_{i,t}, \qquad i = 1, \dots, N,$$
(1)

Our focus is on J_t which has the form

$$d\lambda_{i,t} = \tilde{\alpha}_i (\lambda_{i,\infty} - \lambda_{i,t}) dt + \tilde{\beta}_i dJ_{i,t}, \qquad i = 1, \dots, N,$$
(2)

where the stochastic jump intensity $\lambda_{i,t}$ helps control for the time-varying intensity of the extreme tail shocks at high frequency.

- Boswijk et al. (2018, JoE), Dungey et al. (2018, JoE), Aït-Sahalia et al. (2015, JFE)
- Maheu and McCurdy (2004, JF), Chan and Maheu (2002, JBES)

OUR EXTENSION:

<u>Goal</u>: Systemic (simultaneous) response of assets to specific events <u>Motive</u>: Traders monitoring markets awaiting for FOMC news to assess positions <u>Idea</u>: Going from general (calendar time) form to localized (event time) form <u>Intuition</u>: Analogous to conventional event studies based on high-frequency data

THE LOCALIZED FORM

JUMP-DIFFUSION WITH LOCALIZED (NEWS-BASED) DYNAMICS

LOCALIZED FORM:

Consider the simple localized version of (2):

$$d\lambda_{i,t}^{event} = \tilde{\alpha}_i (\lambda_{i,\infty}^{event} - \lambda_{i,t}^{event}) dt + \tilde{\beta}_i dJ_{i,t}, \qquad i = 1, \dots, N,$$
(3)

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where $\lambda_{i,t}^{event} := [\lambda_{1,t}^{event}, \dots, \lambda_{N,t}^{event}]'$ denotes the stochastic intensities around each FOMC event.

Tail Risk and News Events

- We can use (3) to characterize the dynamics of news-induced shocks that simultaneously hit all assets.
- Systemic tail risk can stem from the common jumping behavior of many stocks, governed by \u03c8_{i+t}^{ivent} in (3), conditional on event times.
- Event times act as reference points for testing/detection.

Schematic Representation

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ESTIMATORS AND TEST STATISTICS

Estimation of λ^{event}

Our variable of interest λ^{event} is latent.

Estimating intensity "before" and "after" news event:

For each stock (i = 1, ..., N) and event (s = 1, ..., S), estimate λ^{event} via

$$\widehat{\lambda}_{i}(k_{n})^{\text{event}} \begin{cases} \widehat{\lambda}_{i}(k_{n})^{\text{pre}} := \frac{\Delta_{n}^{\varpi\widehat{\beta}_{i}}}{k_{n}\Delta_{n}} \sum_{j=1}^{k_{n}} g\left(\frac{|\Delta_{j}^{n}X_{i}^{(\text{pre})}|}{\alpha\Delta_{n}^{\varpi}}\right) \frac{\alpha\widehat{\beta}}{C_{\widehat{\beta}_{i}}(k_{n})} \implies \text{(pre-event)} \\ \widehat{\lambda}_{i}(k_{n})^{\text{post}} := \frac{\Delta_{n}^{\varpi\widehat{\beta}_{i}}}{k_{n}\Delta_{n}} \sum_{j=1}^{k_{n}} g\left(\frac{|\Delta_{j}^{n}X_{i}^{(\text{post})}|}{\alpha\Delta_{n}^{\varpi}}\right) \frac{\alpha\widehat{\beta}}{C_{\widehat{\beta}_{i}}(k_{n})} \implies \text{(post-event)} \end{cases}$$

where $\widehat{\beta}$ controls the vibrancy of sharp fluctuations (jump activity).

 $\begin{array}{l} \longrightarrow \text{ choose } g(\cdot) \\ \longrightarrow \text{ set the window length surrounding events} \\ \longrightarrow \text{ obtain the estimates } \widehat{\lambda}_i(k_n)^{\text{pres}} \text{ and } \widehat{\lambda}_i(k_n)^{\text{post}} \end{array}$

Testing idea:

Does $\hat{\lambda}_i(k_n)^{\text{post}}$ change sharply compared to "benchmark" $\hat{\lambda}_i(k_n)^{\text{pre}}$?

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ESTIMATORS AND TEST STATISTICS

Hypothesis and test statistic

Considering the entire set of events, null and alternative hypotheses:

$$\begin{split} \mathbf{H}_{0}: \quad \boldsymbol{\omega} \in \Omega_{T}^{\mathbf{noSCOJ}} := \Omega_{t}^{\lambda^{\mathbf{event}},0} = \{\boldsymbol{\omega}: \lambda(\boldsymbol{\omega})_{i,t}^{\mathbf{pre}} = \lambda(\boldsymbol{\omega})_{i,t}^{\mathbf{post}}\}, \qquad i = 1, \dots, N, \end{split}$$

$$\mathsf{vs.}$$

$$\begin{split} \mathbf{H}_{a}: \quad \boldsymbol{\omega} \in \Omega_{T}^{\mathbf{SCOJ}} &:= \Omega_{t}^{\lambda^{\mathbf{event}}} &= \{\boldsymbol{\omega}: \lambda(\boldsymbol{\omega})_{i,t}^{\mathbf{pre}} \neq \lambda(\boldsymbol{\omega})_{i,t}^{\mathbf{post}}\}, \qquad i = 1, \dots, N, \end{split}$$

where ω denotes a specific outcome, $\omega\in\Omega.$

Under the null, the event-based test statistic:

Test statistic

$$\mathcal{T}_{i,t}^{(\text{event})} = \sqrt{\frac{k_n \Delta_n}{\Delta_n^{\varpi \widehat{\beta}_i}}} \frac{\widehat{\lambda}_{i,t}^{\text{post}} - \widehat{\lambda}_{i,t}^{\text{pre}}}{\left(\sqrt{\alpha^{\widehat{\beta}_i} C_\beta(2)(\widehat{\lambda}_{i,t}^{\text{post}} + \widehat{\lambda}_{i,t}^{\text{pre}})}\right) \middle/ C_\beta(1)}, \qquad i = 1, \dots, N, \quad (4)$$

which can be computed given $\widehat{\lambda}_{i,t}^{\text{post}}$, $\widehat{\lambda}_{i,t}^{\text{pre}}$, $\widehat{\beta}_i$, k_n , Δ , α and C_{β} .

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LOCALIZED DETECTION WITH EVENT-BASED STEPM Dealing with Spurious Detection

Asymptotic behavior of (4) is not reliable due to multiple testing bias.

We deal with the multiple testing problem in three respects:

- Accounting for the news-induced dependence of test statistics given by (4)
- Asymptotically controlling for the FWE at a given nominal level
- Seeking for high power: better ability to identify false discoveries:
 - ⇒ Eliminate as many "spurious" systemic cojumps as possible
 - \implies Detect as many "real" systemic cojumps as possible

We propose an event-based extension of the stepwise method of Romano and Wolf (2005, *Ecta*).

LOCALIZED DETECTION WITH EVENT-BASED STEPM EVENT-BASED STEPM METHOD

We implement the following procedure to detect event-based systemic cojumps:

Algorithm 1: Event-based StepM

- 1. For each event (define as s = 1, ..., S), use high-frequency data to estimate $\hat{\lambda}_i^{\text{pre}}$ and $\hat{\lambda}_i^{\text{post}}$ for all assets (i = 1, ..., N).
- Compute the test statistic \$\mathcal{T}_{i,t}^{(event)}\$ in Equation (4) conditional on time (t) of each event (s = 1, ..., S) for all assets (i = 1, ..., N). The testing data matrix is \$N \times S\$.
- 3. Relabel the assets (for a given event) in descending order of all $\mathcal{T}_{i,t}^{(\text{event})}$: asset r_1 corresponds to the largest test statistic and asset r_i to the smallest.
- 4. Set j = 1 and $R_0 = 0$ (the number of null hypothesis initially rejected).
- 5. For $R_{(j-1)} + 1 \le i \le N$, if $0 \notin [\mathcal{T}_{r_i,t}^{(event)} \hat{c}_j, \infty)$, reject the null hypothesis $H_0^{(r_i)}$.
- 6. (a) If no (further) null hypotheses are rejected, stop.

(b) Otherwise, denote by R_j the total number of hypotheses rejected so far and, afterward, let j = j + 1. Then, return to step 5.

LOCALIZED DETECTION WITH EVENT-BASED STEPM

EVENT-BASED STEPM METHOD (CONT)

In Algorithm 1, \hat{c}_j denotes the quantiles that we compute directly from the estimated (probability) distribution by using bootstrap.

Under certain assumptions, we have the following result.

Theorem 2: Asymptotic control and consistency

The following statements pertaining to Algorithm 1 are true.

(i) When the null hypothesis is false, the event-based StepM algorithm will reject the null hypothesis with probability 1 as $n \to \infty$.

(ii) The event-based StepM algorithm asymptotically controls the familywise error rate (FWE) at level α ; that is, $\lim_{n} FWE_{\mathbb{P}} \leq \alpha$,

which ensures that bootstrap consistently estimates the limiting distribution of our test statistic.

FINAL TASKS FOR IMPLEMENTATION:

- \implies Use bootstrap to approximate the critical values
- \implies Reject/accept each null for all assets, given the arrival times of events
- \implies The procedure works well: reasonable power and computationally feasible

Monte Carlo Study

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Empirical Analysis

DATA DESCRIPTION

HIGH-FREQUENCY DATA

- A cross-sectional panel of 22 individual DJ stocks and 9 sector ETF indices
- Raw data: tick-by-tick (WRDS) and equally-spaced 15-sec sampling frequency
- Noise correction via price bounceback filtration (Aït-Sahalia et al., 2011, JoE)
- Conventional HF data adjustments and cleaning procedures implemented (excessively low trading activity, missing/constant prices, empty intervals, constant transactions, etc.)
- The sample period: January 31, 2006 through January 30, 2019

FOMC ANNOUNCEMENTS AND MONETARY POLICY (MP) SHOCKS

- Scheduled announcements of the Federal Open Market Committee (FOMC)
- Source: Federal Reserve Bank, Bloomberg and media articles (for cross-check).
- Dates/times of the FOMC news releases (106 events over 2006-2019)
- MP surprise factors: target/path/zero-bound (henceforth Wright factor).
- Construct a revision factor (times when FOMC changes the fed funds target)

MAIN RESULTS

MONETARY POLICY SHOCKS, SYSTEMIC COJUMPS AND CRASHES (DOW JONES)



SEARCHING FOR SYSTEMICALLY IMPORTANT FED EVENTS Dow Jones Stocks

| Date | Rank | SCOJ | Frac_SCOJ | SCRA | Frac_SCRA | Target | Path | Wright |
|------------|------|------|-----------|------|-----------|--------|--------|--------|
| 2008-01-30 | 1 | 22 | 1.00 | 22 | 1.00 | -2.983 | -1.134 | 0.321 |
| 2008-10-29 | 2 | 22 | 1.00 | 22 | 1.00 | -1.859 | -1.816 | -0.383 |
| 2008-12-16 | 3 | 22 | 1.00 | 22 | 1.00 | -3.433 | -3.979 | 3.178 |
| 2009-01-28 | 4 | 22 | 1.00 | 22 | 1.00 | -0.359 | 0.511 | -0.508 |
| 2009-03-18 | 5 | 22 | 1.00 | 22 | 1.00 | 1.060 | -3.904 | 4.991 |
| 2011-08-09 | 6 | 22 | 1.00 | 22 | 1.00 | 0.433 | -0.978 | 1.307 |
| 2018-12-19 | 7 | 22 | 1.00 | 22 | 1.00 | 1.135 | 0.333 | 0.587 |
| 2008-09-16 | 8 | 22 | 1.00 | 21 | 0.95 | 1.802 | 3.438 | -2.443 |
| 2011-09-21 | 9 | 22 | 1.00 | 21 | 0.95 | 0.218 | 1.170 | 0.053 |
| 2009-04-29 | 10 | 22 | 1.00 | 20 | 0.91 | -0.072 | -0.109 | -0.959 |
| 2009-06-24 | 11 | 22 | 1.00 | 20 | 0.91 | -0.106 | 0.431 | -1.735 |
| 2015-09-17 | 12 | 22 | 1.00 | 20 | 0.91 | -1.492 | -1.033 | 0.979 |
| 2007-09-18 | 13 | 22 | 1.00 | 17 | 0.77 | -5.019 | -1.533 | 0.861 |
| 2008-03-18 | 14 | 22 | 1.00 | 17 | 0.77 | 5.490 | 2.195 | -1.252 |
| 2013-06-19 | 15 | 22 | 1.00 | 17 | 0.77 | 0.345 | 0.317 | -2.326 |
| 2009-08-12 | 16 | 22 | 1.00 | 15 | 0.68 | 0.157 | -0.835 | 0.039 |
| 2010-06-23 | 17 | 22 | 1.00 | 13 | 0.59 | -0.012 | 0.794 | 0.138 |
| 2010-08-10 | 18 | 22 | 1.00 | 13 | 0.59 | 0.179 | -0.415 | 0.619 |
| 2009-11-04 | 19 | 22 | 1.00 | 12 | 0.55 | 0.222 | -0.460 | 0.010 |
| 2010-11-03 | 20 | 22 | 1.00 | 9 | 0.41 | 0.219 | -0.318 | -0.211 |

SYSTEMIC EFFECTS OF QUANTITATIVE EASING (QE) QE EVENTS, SYSTEMIC COJUMPS AND CRASHES



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NEWS-DRIVEN REALIZED TAIL RISK

"Fed-Driven" Risk Scores Computed from the Test Statistics (Dow Jones)



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DOES SYSTEMIC TAIL RISK EXPLAIN THE PRE-FOMC DRIFT?

LAGGED REGRESSIONS FOR PRE-FOMC RETURNS

We estimate the following regression model:

$$r_t^{(\text{pre})} = \beta_0 + \beta_x X_{t-1} + \epsilon_t \tag{5}$$

- $r_t^{(pre)}$: Cumulative pre-FOMC high-frequency log-returns
- RS_{t-1} : Lagged realized tail risk scores
- WRS_{t-1} : Lagged weighted realized tail risk scores
- $SCOJ_{t-1}$: Fraction of assets that cojump together at previous meeting's event time

Table: Lagged regressions for pre-FOMC announcement returns

| | | Dependent variable: pre-FOMC announcement returns | | | | | | | | | | |
|---|-----------------------------------|---|----------------------------------|----------------------------------|------------------------------------|------------------------------------|------------------------------------|----------------------------------|--------------------------------|--------------------------------|--------------------------------|--------------------------------|
| | Panel A. All returns | | | | Panel B. Only positive | | | | Panel C. Only negative | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| log(RV(t-1)) | 0.158*** (0.055) | | | | 0.118** (0.046) | | | | 0.041 (0.040) | | | |
| RS(t-1) | | 0.027*** (0.011) | | | | 0.018** (0.009) | | | | 0.009 (0.007) | | |
| WRS(t-1) | | | 0.019** (0.009) | | | | 0.014** (0.007) | | | | 0.005 (0.005) | |
| SCOJ(t-1) | | | | 0.140* (0.080) | | | | 0.081* (0.045) | | | | 0.059 (0.060) |
| Constant | 0.255*** (0.087) | -0.120** (0.054) | -0.062* (0.037) | -0.081 (0.057) | 0.269*** (0.082) | 0.001 (0.034) | 0.035* (0.020) | 0.033 (0.024) | -0.014 (0.059) | -0.121*** (0.044) | -0.096*** (0.028) | -0.114** (0.046) |
| Obs. R ² Res. Std. E. F Statistic | 105 0.079 0.242 8.894*** | 105 0.079 0.242 8.850*** | 105 0.059 0.245 6.401** | 105 0.039 0.247 4.205** | 105 0.131 0.136 15.544*** | 105 0.101 0.138 11.599*** | 105 0.092 0.139 10.418*** | 105 0.039 0.143 4.208** | 105 0.011 0.170 1.182 | 105 0.020 0.170 2.152 | 105 0.010 0.170 0.996 | 105 0.015 0.170 1.577 |

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CONCLUSIONS

OVERVIEW

- A new methodology for quantifying systemic tail risk in a large panel of assets
- High-frequency approach that exploits time-varying jump intensity
- Conditional testing based on the arrival times of news events
- Controlling for spurious detection across assets and limiting data snooping

MAIN TAKEAWAYS AND FINDINGS

- Severe news-induced systemic tail risk that occurs at high frequency
- Systemic market reaction to Fed (FOMC) announcements
- Strong evidence of systemic tail risk in forms of cojumps and crashes
 - \rightarrow <u>cannot</u> be easily diversified away
 - \rightarrow sector rotation strategies are likely to be limited
- Systemic risk helps explain the pre-FOMC drift: significant and sizeable
- Macro news does not create systemic tail risk at high frequency.

FUTURE WORK / EXTENSIONS

- Systemic tail risk, monetary policy news and international portfolio choice
 - \rightarrow A high-frequency approach? \implies Revisiting Das and Uppal (2005, *JF*)?
- Exploring further via market microstructure data (LOB/MBO)? Deep learning for systemic tail risk monitoring?

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SCHEMATIC REPRESENTATION OF THE DETECTION APPROACH



EXTENSIONS AND ROBUSTNESS CHECKS TAILS AND NEWS (I)

Tail risk \Rightarrow probability of extremely large losses

Left tail risk \Rightarrow left skewness of returns

Given (3), how does the tail probability change at high frequency in response to news?

If the news creates a jump with a certain magnitude ξ , we can compute the news-induced tail probability ratio as

$$\frac{\mathbb{P}(|(\lambda_{i,\infty}^{event} + \boldsymbol{\xi})\Delta_n J_i| \ge \alpha \Delta_n^{\varpi})}{\mathbb{P}(|\lambda_{i,\infty}^{event}\Delta_n J_i| \ge \alpha \Delta_n^{\varpi})} \approx \left(\frac{\lambda_{i,\infty}^{event} + \boldsymbol{\xi}}{\lambda_{i,\infty}^{event}}\right)^{\beta_i} = (1 + \boldsymbol{\xi}/\lambda_{i,\infty}^{event})^{\beta_i}, \quad i = 1, \dots, N,$$
(6)

where $\beta_i := [\beta_1, \dots, \beta_N]'$ is the vector of jump activity indices controlling the vibrancy of fluctuations, serving as a tail measure.

This measure is analogous to the estimator of Hill (1975, AoS).

Go back to localized form

TAILS AND NEWS (II)

SIMPLE EXAMPLE:

- For a given value of λ^{event} , if the news generates a jump with large magnitude (e.g., $\xi = 12$), the tail probability ratio is around 1.8.
- In practical terms, this implies that the FOMC event that induces a large jump in each asset will increase the tail probability by 80%.
- Put differently, the likelihood that the investor will be exposed to extreme loss (both left- and right-tails) due to FOMC-driven jumps is now 80% higher than that in the case of no jumps.
- Such a change in the tail probability amplified by news is rather substantial, thereby assets that cojump together pose systemic tail risk.

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MONTE CARLO STUDY

For each stock (i = 1, ..., N), the underlying DGP for log-returns:

$$dX_{i,t} = \sigma_{i,t} dW_{i,t} + \lambda_{\infty} dJ_{i,t}$$
⁽⁷⁾

$$d\sigma_{i,t}^{2} = \kappa(\theta - \sigma_{i,t}^{2}) + \eta \sigma_{i,t}(\phi dW_{i,t} + \sqrt{(1 - \phi^{2})dB_{i,t}}) + \theta \mathbb{1}_{\{S = JT\}}$$
(8)

$$d\lambda_{i,t} = \kappa_{\lambda} (\lambda_{i,\infty} - \lambda_{i,t}) dt + \eta_{\lambda} dB'_{i,t} + \xi \, \mathbb{1}_{\{S = JT\}}, \tag{9}$$

where the vector of Brownian motion $(W_{i,t}, B_{i,t}, B'_{i,t})$ and the vector of β -stable jump processes $J_{i,t}$ are independent from each other.

IMPLEMENTATION

- Select the parameters and calibration values,
- Generate data based on the dynamics given in (8)-(10),
- For each replication, simulate testing time points (representing event times S) that are same for all N asset. Let each asset jump at these pre-determined time points, that is, when jump times (JT) coincide with event times S = JT in (9),
- Determine the pre- and post-event window, (e.g., within hour before and after the event times),
- Compute the estimators, test statistic and apply the StepM detection procedure.

Go back to StepM

MONTE CARLO STUDY: SIMULATION RESULTS (POWER)

| | | | S = 1 | | | S = 10 | |
|-----------|--------|---------|--------|---------|--------|--------|--------|
| Panel A. | | Uncorr | Step-1 | Step-M | Uncorr | Step-1 | Step-M |
| (N = 200) | 15-sec | 99.95% | 99.70% | 99.95% | 99.91% | 95.77% | 99.61% |
| | 1-min | 99.20% | 92.15% | 96.85% | 95.51% | 69.53% | 73.67% |
| Panel B. | | | | | | | |
| (N = 20) | 15-sec | 100.00% | 99.70% | 100.00% | 97.37% | 74.20% | 87.19% |
| | 1-min | 98.10% | 93.30% | 98.20% | 97.87% | 94.40% | 97.54% |
| Panel C. | | | | | | | |
| (N = 20) | 15-sec | 85.70% | 71.20% | 76.10% | 92.86% | 62.52% | 79.70% |
| . , | 1-min | 94.50% | 82.40% | 89.70% | 76.94% | 34.86% | 45.58% |

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Table: Power of the uncorrected and bias-corrected tests based on StepM

"Fed-Driven" Systemic Cojumps and Crashes

| | Events | Assets | Mean | Frac | Stdev | Max | Min | Thr | | |
|---------------------------|--------|--------|-------|------|-------|-----|-----|------|--|--|
| Panel I. Dow Jones stocks | | | | | | | | | | |
| SCOJ | 106 | 22 | 12.77 | 0.58 | 7.76 | 22 | 0 | 0.48 | | |
| SCRA | 106 | 22 | 5.75 | 0.26 | 7.61 | 22 | 0 | 0.20 | | |
| | | | | | | | | | | |
| Panel II. Sector ETFs | | | | | | | | | | |
| SCOJ | 106 | 9 | 4.93 | 0.55 | 3.51 | 9 | 0 | 0.50 | | |
| SCRA | 106 | 9 | 2.08 | 0.23 | 2.91 | 9 | 0 | 0.18 | | |

MONETARY POLICY SHOCKS, SYSTEMIC COJUMPS AND CRASHES (SECTOR ETFS)



IN SEARCH OF SYSTEMICALLY IMPORTANT EVENTS (SECTOR ETFs)

| Date | Rank | SCOJ | Frac_SCOJ | SCRA | Frac_SCRA | Target | Path | Wright |
|------------|------|------|-----------|------|-----------|--------|--------|--------|
| 2007-12-11 | 1 | 9 | 1.00 | 9 | 1.00 | 1.459 | -1.898 | 2.164 |
| 2011-08-09 | 2 | 9 | 1.00 | 9 | 1.00 | 0.433 | -0.978 | 1.307 |
| 2018-03-21 | 3 | 9 | 1.00 | 9 | 1.00 | 0.391 | -0.259 | -0.006 |
| 2018-12-19 | 4 | 9 | 1.00 | 9 | 1.00 | 1.135 | 0.333 | 0.587 |
| 2008-01-30 | 5 | 9 | 1.00 | 8 | 0.89 | -2.983 | -1.134 | 0.321 |
| 2011-09-21 | 6 | 9 | 1.00 | 8 | 0.89 | 0.218 | 1.170 | 0.053 |
| 2008-03-18 | 7 | 9 | 1.00 | 7 | 0.78 | 5.490 | 2.195 | -1.252 |
| 2008-12-16 | 8 | 9 | 1.00 | 7 | 0.78 | -3.433 | -3.979 | 3.178 |
| 2009-03-18 | 9 | 9 | 1.00 | 7 | 0.78 | 1.060 | -3.904 | 4.991 |
| 2009-11-04 | 10 | 9 | 1.00 | 7 | 0.78 | 0.222 | -0.460 | 0.010 |
| 2016-01-27 | 11 | 9 | 1.00 | 7 | 0.78 | -0.013 | 0.163 | 0.462 |
| 2007-08-07 | 12 | 9 | 1.00 | 6 | 0.67 | 1.361 | 0.414 | -0.646 |
| 2008-10-29 | 13 | 9 | 1.00 | 6 | 0.67 | -1.859 | -1.816 | -0.383 |
| 2009-01-28 | 14 | 9 | 1.00 | 6 | 0.67 | -0.359 | 0.511 | -0.508 |
| 2009-06-24 | 15 | 9 | 1.00 | 6 | 0.67 | -0.106 | 0.431 | -1.735 |
| 2010-11-03 | 16 | 9 | 1.00 | 6 | 0.67 | 0.219 | -0.318 | -0.211 |
| 2009-04-29 | 17 | 9 | 1.00 | 5 | 0.56 | -0.072 | -0.109 | -0.959 |
| 2015-09-17 | 18 | 9 | 1.00 | 5 | 0.56 | -1.492 | -1.033 | 0.979 |
| 2015-12-16 | 19 | 9 | 1.00 | 4 | 0.44 | 0.574 | 0.465 | -0.399 |
| 2008-04-30 | 20 | 9 | 1.00 | 3 | 0.33 | -1.652 | -1.239 | 1.182 |

EXTENSIONS AND ROBUSTNESS CHECKS QE EVENTS, SYSTEMIC COJUMPS AND CRASHES

| Data | T | T | Desiring (see | 60011 | 14/2012 | 6601 | 6601 | CCDA | CCDA | D1/ | 0.01/ |
|-------------------|-------|------|--|-------|---------|-----------|------|-----------|------|-------|-----------|
| Dates 2009112E | 0.1E | OE1 | Turns of sumt: EOMC statement. Europeing of OE | G2011 | VV2012 | SCOJ-stat | SCOJ | SCRA-stat | SCRA | NA NA | DRV ratio |
| 20081125 | 0.15 | QLI | Initial I SAP appouncement. The Fed appouncer purchases of | -22 | 0.75 | 1975 | 1974 | 1974 | 1974 | 1974 | 1974 |
| | | | \$100 billion in GSE debt and up to 500 billion in MBS. Creation | | | | | | | | |
| | | | of the Term Asset-Backed Security Loan Facility (TALE) | | | | | | | | |
| 20081201 | 13:45 | OF1 | Type of event: Bernanke Speech - Expansion of QE Chair- | -19 | 0.84 | 5.24 | 22 | 4.06 | 20 | 1 4 1 | 1 44 |
| | | | man Bernanke mentions that the Fed could purchase long-term | | | | | | | | |
| | | | Treasuries. | | | | | | | | |
| 20081216 | 14:15 | QE1 | Type of event: FOMC statement - Expansion of QE. | -26 | 2.22 | (11.80) | 22 | (7.92) | 22 | 2.78 | 2.86 |
| | | | The FOMC "evaluates" the potential benefits of purchasing | | | | | | | | |
| | | | longer-term Treasury securities. Also FED funds target rate | | | | | | | | |
| | | | reduced to the range 0- 0.25 | | | | | | | | |
| 20090128 | 14:15 | QE1 | Type of event: FOMC statement – Expansion of QE. | 14 | -0.23 | 8.35 | 22 | 6.09 | 22 | 2.27 | 2.45 |
| | | | The Fed is ready to expand agency debt and MISS purchases, | | | | | | | | |
| 20000219 | 14.15 | OE1 | as well as to purchase long term treasuries. | 47 | 2.41 | (11.05) | 22 | (7.79) | 22 | 2.52 | 2.45 |
| 20090310 | 14.15 | 4 | The Fed will purchase an additional \$750 billion in agency MBS | -41 | 3.41 | (11.05) | | (1.10) | | 2.33 | 2.45 |
| | | | and an additional \$100 billion in Agency Debt. Moreover, the | | | | | | | | |
| | | | FOMC decided to purchase up to \$300 billion of longer-term | | | | | | | | |
| | | | Treasury securities over the following six months. | | | | | | | | |
| 20090812 | 14:15 | QE1 | Type of event: FOMC statement - Phase out of QE. | 5 | 0.15 | 5.57 | 22 | 4.23 | 21 | 2.11 | 2.30 |
| | | | The Fed will slow the pace of the LSAP by purchasing the full | | | | | | | | |
| | | | amount by the end of October instead of mid- September. | | | | | | | | |
| 20090923 | 14:15 | QE1 | Type of event: FOMC statement - Phase out of QE. | -3 | 0.85 | 5.14 | 21 | 3.73 | 20 | 2.23 | 2.26 |
| | | | The Fed will slow the purchases of agency MBS and agency | | | | | | | | |
| | | | debt, finishing the purchases by the end of 2010Q1. Treasury | | | | | | | | |
| 20001104 | 14.15 | OE1 | Turns of summer EOMC statements. Phase out of OE | 6 | 0.12 | 6.50 | 22 | 4 20 | 20 | 2.40 | 3.37 |
| 20091104 | 14.15 | QLI | The amount of agency debt will be halted at \$175 billion in- | 0 | 0.12 | 0.52 | 22 | 4.30 | 20 | 2.40 | 2.21 |
| | | | stead of \$200 billion. | | | | | | | | |
| 20100810 | 14:15 | QE2 | Type of event: FOMC statement - Expansion of QE. | NA | 0.57 | 5.85 | 22 | 4.14 | 22 | 2.62 | 2.55 |
| | | | The Fed will reinvest principal payments from agency debt and | | | | | | | | |
| | | | agency mortgage-backed securities in longer-term Treasury se- | | | | | | | | |
| | | | curities. Holdings of Treasury securities will be rolled over as | | | | | | | | |
| | | | they mature. | | | | | | | | |
| 20100827 | 10:00 | QE2 | Type of event: Bernanke speech – Expansion of QE. | NA | -0.83 | NA | NA | NA | NA | NA | NA |
| | | | bernanke mentions potential policy options for further easing, | | | | | | | | |
| 20101015 | 14-15 | OE2 | Type of event: Bernanke speech – Evpansion of OE | NA | -0.21 | [1 88] | 8 | [1.01] | 3 | 1 32 | 1.26 |
| 20101010 | 14.15 | der. | The Fed is prepared to provide additional accommodation if | 104 | -0.21 | [1:00] | 0 | [1:01] | 5 | 1 | 1.10 |
| | | | needed to support the economic recovery. | | | | | | | | |
| 20101103 | 14:15 | QE2 | Type of event: FOMC statement - Expansion of QE. | NA | -0.05 | 5.55 | 22 | 3.81 | 22 | 2.37 | 2.11 |
| | | | The Fed will purchase a further \$600 billion of longer-term Trea- | | | | | | | | |
| | | | sury securities by the end of the second quarter of 2011, a pace | | | | | | | | |
| | | | of about \$75 billion per month. | | | | | | | | |

"Fed-Driven" Risk Scores Computed from the Test Statistics (Sector ETFs)



Fed-driven stress scores (sector ETFs)

NEGATIVE VERSUS POSITIVE MONETARY POLICY SHOCKS (DOW JONES)



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MACROECONOMIC NEWS ANNOUNCEMENTS

| News type | News ticker | Release time | Frequency | Relevance | Events |
|--------------------|-------------|--------------|-----------|-----------|--------|
| GDP Annualized QoQ | CQOQ | 8:30 | Quarterly | 96.81 | 54 |
| Unemployment Rate | USURTOT | 8:30 | Monthly | 89.28 | 163 |
| CPI MoM | CPI CHNG | 8:30 | Monthly | 95.41 | 163 |
| ISM Manufacturing | NAPMPMI | 10:00 | Monthly | 95.83 | 163 |
| New Home Sales | NHSLTOT | 10:00 | Monthly | 90.44 | 163 |

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Systemic Reaction Heat Maps for News Announcements



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Systemic Reaction Heat Maps for News Announcements (Cont)



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MARKET MICROSTRUCTURE NOISE

"Why not using pre-averaging method to remove noise?"

- Pre-averaging method has been widely shown as a very effective tool in eliminating the impact of noise in the estimation of realized quantities (and jump detection). → e.g. Hautsch and Podolskij (2013, *JBES*), Podolskij and Vetter (2009, *Bernoulli*)
- Despite its viable usefulness for noise reduction, pre-averaging method might be of limited help in our context for estimating jump compensator related quantities → Bücher and Vetter (2013, AoS).
- Recall our variable of interest λ (stochastic jump intensity) which belongs to Lévy measure.
- The challenge: Bücher and Vetter (2013, AoS) show pre-averaging does not yield a consistent estimator of the tail of a Lévy measure.
- See Bücher and Vetter (2013, AoS) and Boswijk et al. (2018, JoE) for the discussion.
- We instead use price bounceback filtration (as in Aït-Sahalia et al., 2011, JoE) to eliminate the potential impact of noise.

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